

Major Exam

MAL2040: Design and Analysis of Experiments

Date: 2026-04-27

Duration: 180 minutes

Total: 40 points

Name: _____

Roll No.: _____

Instructions:

- Hand calculators are allowed.
- State all assumptions used for full credit.
- Unnecessary explanations can result in loss of credit.

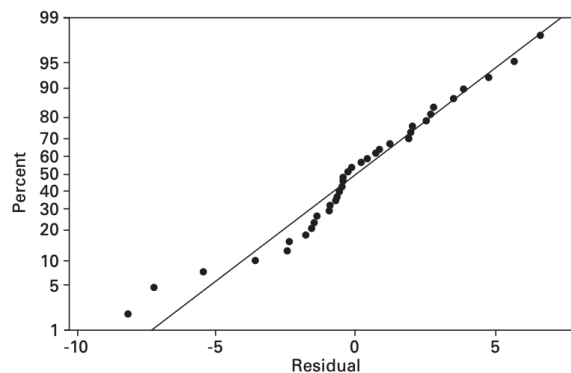
Problem 1. [2 points]

State whether each of the following statements is True or False. No justification is required.

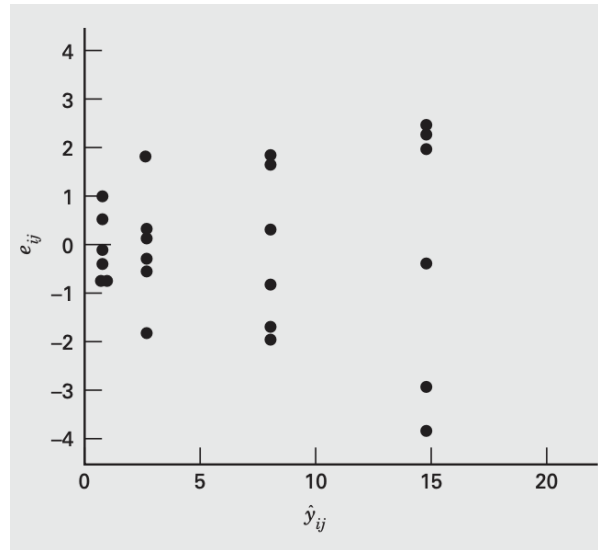
- (i) If the block effect is significant in an RCBD, then the treatment F -test is more likely to detect a true treatment effect compared to in a completely randomized design.
- (ii) In a $p \times p$ Latin square design, removing one row still allows the treatment effects to be estimated independently of the row and column effects.
- (iii) In the intrablock analysis of a BIBD, the adjusted treatment totals Q_i satisfy $\sum_{i=1}^a Q_i = 0$.
- (iv) In response surface methodology, the method of steepest ascent is used to explore the region near a suspected optimum where a second-order model is to be fitted.

Problem 2. A textile manufacturer studies the effect of four dye formulations (1, 2, 3, 4) on the colourfastness of a fabric. Since fabric bolts may differ in quality, a randomized complete block design is used with five bolts as blocks. After fitting the RCBD model, the following diagnostic plots are obtained.

- (a) Write down the statistical model for this experiment, clearly defining every term and parameter, and state all assumptions. [2 points]
- (b) State the null and alternative hypotheses for the treatment F -test, and write down the F -statistic and its degrees of freedom. [1 point]
- (c) The figure below is a normal probability plot of the residuals. What conclusions do you draw? [1 point]



- (d) The figure below shows residuals plotted against fitted values. What conclusions do you draw? [1 point]



Problem 3. An experiment is conducted to compare the yields (in %) of four chemical formulations (A, B, C, D). Since batch of raw material and operator may introduce variability, a 4×4 Latin square design is used with batches as rows and operators as columns. The observed yields, along with row and column totals, are:

	Op. 1	Op. 2	Op. 3	Op. 4	Row Total
Batch 1	$A = 21$	$B = 17$	$C = 26$	$D = 19$	83
Batch 2	$B = 14$	$C = 21$	$D = 17$	$A = 25$	77
Batch 3	$C = 19$	$D = 16$	$A = 21$	$B = 15$	71
Batch 4	$D = 18$	$A = 22$	$B = 14$	$C = 21$	75
Col. Total	72	76	78	80	306

You are given that $\sum_{\text{all obs}} y_{ijk}^2 = 6046$.

- Compute the correction factor $CF = y_{...}^2/N$ and SS_{Total} . [1 point]
- Compute SS_{Batches} and $SS_{\text{Operators}}$. [1 point]
- Compute the treatment totals for A, B, C, D , and hence compute $SS_{\text{Treatments}}$. [1 point]
- Compute SS_{Error} . [1 point]
- Construct the ANOVA table. [1 point]
- Test whether there is a significant difference in yield among the four formulations at $\alpha = 0.05$. The critical value is $F_{0.05,3,6} = 4.76$. [1 point]

Problem 4.

- (a) A BIBD is proposed with $a = 6$ treatments, block size $k = 4$, and each treatment replicated $r = 5$ times.

Compute λ and b . Find a BIBD with these parameters if it exists, or show that such a BIBD cannot exist. **[2 points]**

- (b) Consider a BIBD with a treatments and b blocks of size k , in which each treatment appears in r blocks and each pair of treatments appears together in λ blocks. The model is:

$$y_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}, \quad \varepsilon_{ij} \sim \text{NID}(0, \sigma^2).$$

The adjusted treatment total for treatment i is defined as

$$Q_i = y_{i.} - \frac{1}{k} \sum_{j=1}^b n_{ij} y_{.j},$$

where $n_{ij} = 1$ if treatment i appears in block j , and 0 otherwise.

- (i) Write down the normal equations by minimising the error sum of squares:

$$L = \sum_{i=1}^a \sum_{j=1}^b n_{ij} (y_{ij} - \mu - \tau_i - \beta_j)^2.$$

[2 points]

- (ii) Apply the usual constraints and solve the resulting system of linear equations to find a formula for the intrablock estimator $\hat{\tau}_i$ of the treatment effect τ_i in terms of the adjusted treatment totals Q_i . **[2 points]**

- (c) In the interblock analysis of the BIBD, the block effects β_j are assumed to behave like random noise. The model is:

$$y_{.j} = k\mu + \sum_{i=1}^a n_{ij}\tau_i + \text{error}, \quad j = 1, \dots, b,$$

where the error term includes the block effects.

- (i) Write down the normal equations by minimising the error sum of squares:

$$L = \sum_{j=1}^b \left(y_{.j} - k\mu - \sum_{i=1}^a n_{ij}\tau_i \right)^2.$$

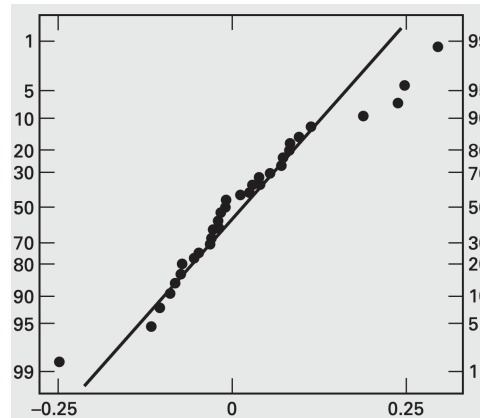
[2 points]

- (ii) Apply the usual constraint and solve the resulting system of linear equations to find a formula for the interblock estimator $\tilde{\mu}$ of the grand mean μ . **[2 points]**

Problem 5.

- (a) In a 2^k factorial design, state the total number of (i) treatment combinations, (ii) main effects, and (iii) two-factor interaction effects. Give your answers in terms of k . **[1 point]**
- (b) State the *sparsity of effects principle*. Explain how it is used to identify estimate the error in an unreplicated 2^k design. **[1 point]**

- (c) A 2^4 factorial experiment is run with factors A, B, C, D . After analysis, it is found that B and all interactions involving B are negligible. What does the original 2^4 design reduce to, and how many effective replicates does the experiment have? [1 point]
- (d) The figure below shows a normal probability plot of the estimated effects from an unrepliated 2^4 experiment. How many effects appear to be active? Justify your answer. [2 points]

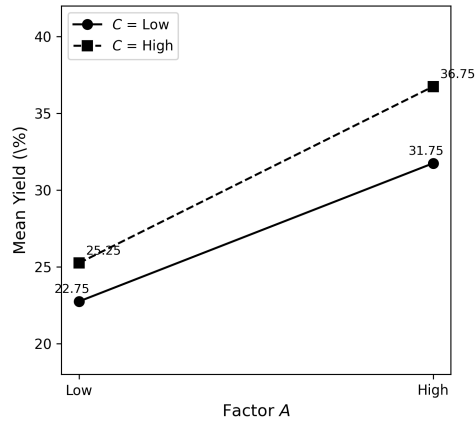


Problem 6. A chemical engineer studies three process factors: reaction temperature (A), pressure (B), and catalyst concentration (C), each at two levels. A 2^3 factorial experiment is run with $n = 2$ replicates. The observed yields (in %) are:

Treatment	Replicate 1	Replicate 2	Treatment Total
(1)	28	25	53
a	36	32	68
b	18	20	38
ab	30	29	59
c	31	28	59
ac	40	37	77
bc	20	22	42
abc	34	36	70

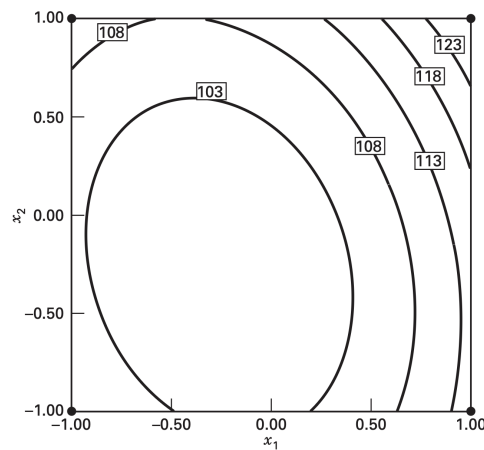
You are given: $\sum_{\text{all obs}} y^2 = 14244$, grand total $y_{\dots} = 466$, $CF = y_{\dots}^2/N = 13572.25$.

- (a) Compute SS_{Total} . [1 point]
- (b) Write down the sign table for a 2^3 factorial design in standard order. [1 point]
- (c) Compute the contrast and sum of squares for each of the seven effects (A, B, AB, C, AC, BC, ABC). [2 points]
- (d) Compute SS_{Error} and construct the complete ANOVA table. [1 point]
- (e) Using $F_{0.05,1,8} = 5.32$, state which effects are significant. What factor-level combination maximises yield? [1 point]
- (f) The figure below shows an interaction plot for factors A and C (averaged over the levels of B). Interpret this plot. Does it agree with your conclusion from part (d)? [1 point]



Problem 7.

- (a) The contour plot below shows the fitted response surface $\hat{y}(x_1, x_2)$ from a second-order model. Based on the plot, determine whether the stationary point of this surface is a local maximum, local minimum, or saddle point. Justify your answer by referring to specific features of the plot. **[2 points]**



- (b) A first-order response surface model is fitted in two coded variables, giving

$$\hat{y} = 52.3 + 1.8x_1 + 0.9x_2,$$

where $x_1 = (\xi_1 - 25)/5$ and $x_2 = (\xi_2 - 80)/10$, with ξ_1 = reaction time (minutes) and ξ_2 = reaction temperature ($^{\circ}\text{C}$).

- State the direction of steepest ascent in the coded variable space.
 - Taking a step size of $\Delta\xi_1 = 5$ minutes, compute $\Delta\xi_2$.
 - Starting from $(\xi_1, \xi_2) = (25, 80)$, tabulate the first three points along the path of steepest ascent. **[2 points]**
- (c) After moving along the steepest ascent path, a second-order model is fitted. The stationary point is found at $\mathbf{x}_0 = (1.2, -0.8)^T$, and the eigenvalues of the matrix of second-order coefficients are $\lambda_1 = -3.2$ and $\lambda_2 = -1.5$. What type of stationary point is this? Compute the optimal operating conditions in the natural variables ξ_1 and ξ_2 . **[1 point]**

End of Exam